1 Random Walk

1.1 Symmetric Random Walk

1.1.1 Definition

Starting from coin toss, def:

$$\begin{aligned} X_j &= 1 & \text{if } w_j &= H \\ &= -1 & \text{if } w_j &= T \end{aligned}$$

Let $M_k = \sum_{j=1}^k X_j$, M_k is symmetric random walk, and its distribution is binomial.

1.1.2 Expectation and Variance

 $E(X_j) = 0, Var(X_j) = 1,$ $Var(M_k - M_l) = \sum_{i=l}^k Var(X_i) = k - l;$

1.1.3 Martingale Property

 $E(M_{l}|F_{k}) = E((M_{l} - M_{k} + M_{k})|F_{k}) = E((M_{l} - M_{k})|F_{k}) + E(M_{k}|F_{k}) = 0 + M_{k} = M_{k}$

1.2 Scaled Random Walk

1.2.1 Scaled Symmetric Random Walk Definition

$$W^{(n)}(t) = 1/\sqrt{n}M_{nt}$$

= $1/\sqrt{n}\sum_{i=1}^{nt}X^{(n)}(t)$
= $\sum_{i=1}^{nt}1/\sqrt{n}X^{(n)}(t)$

1.2.2 Expectation and Variance

$$E(W^{(n)}(t) - W^{(n)}(s)) = 0$$

$$Var(W^{(n)}(t) - W^{(n)}(s)) = n(t-s)Var(1/\sqrt{n}X^{(n)}(t-s)) = t-s$$

1.3 Brownian Motion

1.3.1 Limit of Scaled Symmetric Random Walk: Brownian Motion

When $n - > \inf$, based on Central Limit Theorem $W^{(n)}(t)$ starting from 0 follows a normal distribution N(0, t) We view $W^{(n)}(t)$ as a Brownian motion

1.3.2 Quadratic variation of Brownian motion

Given a Brownian motion W_{t_j} Define

$$Q = \sum_{j=0}^{n-1} (W_{t_{j+1}} - W_{t_j})^2$$
$$E(Q) = \sum_{j=0}^{n-1} E((W_{t_{j+1}} - W_{t_j})^2)$$
$$= \sum_{j=0}^{n-1} (t_{j+1} - t_j)$$
$$= T$$

$$Var(Q) = \sum_{j=0}^{n-1} 2(t_{j+1} - t_j)^2$$

<= 2|C|T(where||C|| = max|t_{j+1} - t_j|)
When|C| = 0,
Var(Q) = 0

This is simply written in differential form dWdW = dt

1.3.3 From Scaled Asymmetric Random Walk to Geometric Brownian motion

Consider a scaled asymmetric random walk with factor of σ and drift α

$$W^{(n)}(t) = \sigma(\frac{1}{\sqrt{n}}M_{nt}) + \alpha t$$

Construct new random variable which satisfies $\delta S/S = W^{(n)}(t)$, which is equivalent to

$$S + = S(1 + \alpha/n + \sigma/\sqrt{n})$$
$$S - = S(1 + \alpha/n - \sigma/\sqrt{n})$$

Assuming number of heads is H_{nt} , number of tails is T_{nt} Then

$$S_n(t) = S_n(0)(1 + \alpha/n + \sigma/\sqrt{n})^{(H_{nt})}(1 + \alpha/n - \sigma/\sqrt{n})^{(T_{nt})}$$

$$\begin{split} log(S_n(t)) &= log(S_n(0)) + (H_{nt})log(1 + \alpha/n + \sigma/\sqrt{n}) + (T_{nt})log(1 + \alpha/n - \sigma/\sqrt{n}) \\ &= log(S_n(0)) + (H_{nt})(\alpha/n + \sigma/\sqrt{n} - 1/2(\alpha/n + \sigma/\sqrt{n})^2) \\ &+ (T_{nt})(\alpha/n - \sigma/\sqrt{n} - 1/2(\alpha/n - \sigma/\sqrt{n})^2) \\ &= log(S_n(0)) + (H_{nt} + T_{nt})(\alpha/n - \sigma^2/n + O(n^{-3/2})) + (H_{nt} - T_{nt})(\sigma/\sqrt{n}) \\ &= log(S_n(0)) + nt(\alpha/n - 1/2\sigma^2/n) + M_{nt}\sigma/\sqrt{n} \\ &= log(S_n(0)) + (\alpha - \frac{1}{2}\sigma^2)t + W^n(t)\sigma \end{split}$$

Then we can prove when $n - > \infty$ The distribution of $S_n(t)$ converges to the distribution of

$$S(t) = S(0)exp((\alpha - \frac{1}{2}\sigma^2)t + \sigma W(t))$$

Where W(t) is a normal random variable with mean 0 and variance t.